

Insight into Binary Star Formation via Modelling Visual Binaries Datasets

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Abstract. We describe the project aimed at finding initial distributions of binary stars over masses of components, mass ratios of them, semi-major axes and eccentricities of orbit, and also pairing scenarios by means of Monte-Carlo modeling of the sample of about 1000 visual binaries of luminosity class V with Gaia DR1 TGAS trigonometric parallax larger than 2 mas, limited by $2 \leq \rho \leq 200$ arcsec, $V_1 \leq 9.5^m$, $V_2 \leq 11.5^m$, $\Delta V \leq 4^m$, which can be considered as free of observational incompleteness effects. We present some preliminary results which allow already to reject initial distributions of binaries over semi-major axes of the orbits more steep than $\propto a^{-1.5}$.

Keywords: binary stars, stellar formation, modeling

1 Introduction

Majority of stars accessible for detailed observational study appear to be binary ones. Interaction between binary star components in the course of their evolution results in a rich variety of astrophysical phenomena and objects. Study of the structure and evolution of binary stars is one of the most actively developing fields of the modern astrophysics.

Among fundamental problems aimed by these studies is the one of initial distributions of binary stars over their main parameters:

- mass of the primary component M_1 ,
- mass ratio of components $q = M_2/M_1$,
- and semi-major axis a of component orbits,

which, combined, we will call “the birth function” (henceforth, BF).

Most important, BF is, first, a benchmark for the theories of star formation and, second, the base for the estimates of the number of objects in the models of different stellar populations and model rates of various events, e. g., supernovae explosions etc.

In the present study, we assume that BF is defined by three fundamental functions describing distribution of stars over initial mass of primary component M_1 , mass ratio of components q , and semi-major axes of orbits a [25]. It was suggested by Vereshchagin et al. (1988) [35] that BF for visual binaries has the form

$$d^3 N \propto M_1^{-2.5} dM_1 \cdot d \log a \cdot q^{-2.5} dq \text{ year}^{-1} \quad (1)$$

where M_1 and a are expressed in solar units. As a “minor” characteristics, we consider eccentricity of orbits e . The aim of the current paper is presentation of preliminary results of the assessment of BF by means of

comparison of results of Monte-Carlo model of the local population of field visual binaries with their observed sample.

We probe, for a given type of stars, whether the synthetic dataset differs significantly due to the change of initial fundamental distributions, and how the change of every distribution affects it. For this purpose, we compare synthetic populations for different pairing functions and particular sets of fundamental functions. We attempt to find whether certain initial distributions or combinations of initial distributions result in synthetic datasets incompatible with observational data at certain significance level and, on the contrary, whether certain initial distributions or combinations of initial distributions provide synthetic dataset best compatible with observational data, hopefully, at certain significance level.

The model also accounts for star formation rate, stellar evolution and takes into account observational selection effects. The model is compared to the dataset compiled as described by Kovaleva et al [18] with addition of the data on parallaxes from Gaia DR1 TGAS [9].

Besides, our model allows to obtain estimates for the fraction of binary stars that remains unseen for different reasons and is observed as single objects and to investigate how these fractions depend on the initial distributions of parameters. Such estimates are important, for instance, as an approach toward recovering actual multiplicity fraction, mass hidden in binaries, as well as toward models of different stellar populations.

The model and observational data are described in chapters 2 and 3, respectively. Some considerations on the choice of theoretical models are described in chapter 4. Results and conclusions are presented in chapter 5. In chapter 6 we outline the plans of future studies.

2 The model

Visual binaries are observed, mostly, in the immediate solar vicinity. Therefore, we consider them to be distributed up to the distance of 500 pc in radial direction and according to a barometric function along z . The scale height z for the stars of different spectral types and, respectively, masses was studied, e. g., in [3,10,12,20,31]. Synthesizing results of these studies, we assume $|z| = 340$ pc for low-mass ($\leq 1 M_{\odot}$) stars, 50 pc for high-mass ($\geq 10 M_{\odot}$) stars, and linear $|z| - \log M$ relation for intermediate masses.

For such a small volume we can neglect the radial gradient (Huang et al. 2015) [14]. We also ignore interstellar extinction.

To simulate stellar pairs we use different pairing functions (scenarios), mostly taken from Kouwenhoven's list [16].

It includes random pairing and other scenarios, where two of the four parameters (primary mass, secondary mass, total mass of the system, mass ratio) are randomized, and other are calculated. Table 1 contains

the short summary of the used pairing functions.

Masses of the components or total masses of the binaries were drawn randomly from Salpeter [32] or Kroupa [21] initial mass functions (IMF), separation a was drawn from one of the following distributions: $\propto a^{-1}$, $\propto a^{-1.5}$, $\propto a^{-2}$, and eccentricity e was distributed assuming following options: (i) all orbits are circular, (ii) eccentricities obey thermal distribution $f_e(e) = 2e$, and (iii) equiprobable distribution $f_e(e) = 1$. We adopt random orbit orientation. Mass ratio q , when needed, is randomly drawn from $\propto q^{\beta}$ distribution, where β is adopted to be -0.5 , 0 or -0.5 . The lower limit for q is determined by mass limits $[0.08 \dots 100] M_{\odot}$. Certain pairing functions, such as RP, PCR, PSCP and TPP, do not allow independent random distribution of mass ratios, it is calculated from masses of components.

Table 2 contains short summary on initial distributions used in the modelling. Some cells are empty because the pre-planned distributions are not implemented as yet. The total number of possible combinations of initial distributions considered as yet is 144, equal to the number of possible combinations of s , m , q , a , e in Table 2 and regarding that $s0$ and $s5$ scenarios do not imply independent initial distribution over q (see Table 1). Any combination of distributions listed in Table 2 can be conveniently referred, for instance, as "s2m0q5a1e0".

To account for star formation rate we adopt $SFR(t) = 15 e^{-t/7}$, where the time t is expressed in Gyr (Yu & Jeffery 2010 [36]). Disc age is assumed to be equal to 14 Gyr.

Currently, we consider the following stellar evolutionary stages: MS-star, red giant, white dwarf, neutron star, black hole. The objects in the two latter stages do not produce visual binaries (though they contribute to the statistics of pairs, observed as single stars, see Section 5.2 below). We do not consider brown dwarfs and pre-MS stars here, as they are extremely rarely observed among visual binaries and their multiplicity rate is substantially lower than for more massive stars (Allers 2012) [1]. As we deal with wide pairs only, we assume the components to evolve independently. To calculate evolution of stars and their observational properties we used analytical expressions derived by Hurley et al [15] and assumed solar metallicity for all generated stars.

To normalize the number of simulated objects, we use estimates of stellar density in the solar neighborhood, based on recent Gaia results [4]. The data for A0V-K4V stars presented by Bovy (2017) [4] give 0.01033 stars per pc^3 . This means that in the 500 pc sphere we generate about 43300 pairs of stars.

For the generated objects, we determine observational parameters, in particular, the brightness of components, their evolutionary stage and projected separation. Then we apply a filter to select a sample of stars, which can be compared with observational data (see the next section).

Table 1 Summary of considered pairing functions (scenarios)

Abbreviation	Full name	Scheme
RP	Random Pairing	rand($M_1, M_2, [M_{min} \dots M_{max}]$); sort(M_1, M_2); calc(q).
PCRP	Primary Constrained Random Pairing	rand($M_1, [M_{min} \dots M_{max}]$); rand($M_2, [M_{min} \dots M_{max}], M_1 = const$ until $M_2 < M_1$); calc(q).
PCP	Primary Constrained Pairing	rand(M_1, q); calc(M_2).
SCP	Split-Core Pairing	rand($M_{tot}, [2M_{min} \dots 2M_{max}]$); rand(q); calc(M_1, M_2).
PSCP	Primary Split-Core Pairing	rand($M_{tot}, [2M_{min} \dots 2M_{max}]$); rand($M_1, [0.5(M_1 + M_2) \dots M_{max}]$, until $M_1 < M_{tot}$); calc(M_2); calc(q).
TPP	Total Primary Pairing	rand($M_{tot}, [2M_{min} \dots 2M_{max}]$); rand($M_1, [M_{min} \dots M_{max}]$, until $M_1 < M_{tot}$); calc(M_2); sort(M_1, M_2); calc(q).

Note: M_{tot} , M_1 , M_2 – total mass of the binary, primary mass and secondary mass, respectively; M_{min} , M_{max} – lower ($0.08 M_{\odot}$) and upper ($100 M_{\odot}$) limits set for masses; $q = M_2/M_1$ – mass ratio. The meaning of abbreviations is the following: “rand” – randomizing, “calc” – calculation, “sort” – sorting.

Table 2 Summary of applied initial distributions

sN	Scenario (s)	mN	IMF (m)	qN	Mass ratio (q)	aN	Semi-major axis (a)	eN	Eccentricity (e)
0	RP	0	Salpeter	0	flat, $f = 1$	0	power, $f \sim a^{-1}$	0	thermal, $f = 2e$
		1	Kroupa			1	power, $f \sim a^{-1.5}$	1	delta, $f = \delta(0)$
2	PCP					2	power, $f \sim a^{-2}$	2	flat, $f = 1$
3	SCP								
				4	power, $f \sim q^{-0.5}$				
5	TPP			5	power, $f \sim q^{0.5}$				

3 Observational data for comparison

To compare our simulations with observational data, we use the most comprehensive list of visual binaries WCT [17], compiled on the base of the largest original catalogues WDS [24], CCDM [5] and TDSC [8]. These data were refined or corrected for mistaken data, optical pairs, effects of higher degrees of multiplicity, sorted by luminosity class (primarily, to select pairs with both components on the main-sequence), and appended by parallaxes. A refined dataset for comparison was selected from the data, so as to avoid regions of observational

incompleteness in the space of observational parameters. The procedure of dataset compilation and analysis described in details in [17,18] was improved due to use of new trigonometric parallaxes from TGAS DR1 Gaia [9] that allowed to re-obtain constraints to avoid regions of observational incompleteness.

Out of simulated objects we select pairs, satisfying the same observational constraints, as the refined observational set does, namely: projected separation $2 < \rho < 200$ arcsec, primary component visual magnitude $V_1 < 9.5^m$, secondary component visual magnitude $V_2 < 11.5^m$, magnitude difference $\Delta V \equiv |V_2 - V_1| \leq 4^m$ (henceforth, “synthetic dataset”). For the purposes of

correct comparison, we also limit refined set of observational data by 500~pc distance.

We construct distributions of synthetic datasets over the following parameters: primary and secondary magnitude, magnitude difference, projected separation, parallax.

Then we compare the synthetic distributions with refined observational ones using χ^2 two-sample test. We deem, the better result of comparison, the closer our assumptions on pairing scenarios, initial distributions of masses, mass ratio, separation and eccentricity are to reality. The refined set of observational data contains $N = 1089$ stars. To compare them properly with results of our simulations we need to use histograms with $n = 5 \log N$ bins [33], i. e., 15 ones.

4 Some reflections concerning selection of models

In the selection of trial initial distributions for the model we adopted the following approach: we started with well established or widely used in the literature functions for $f(M)$, $f(a)$, $f(e)$ and then stepped aside from them to test, whether the algorithm would be able to feel difference at all. We preferred simple analytical expressions, supposing we would pass to more complicated ones later if we find it necessary.

Thus, we use traditional Salpeter's IMF [32] along with the much more recent and generally accepted Kroupa's one [21]. In spite of the statement by Duchêne and Kraus [7] that no observed dataset agrees with random pairing scenario, we use the latter among other ones.

On the other hand, for semi-major axis distribution we applied as yet only commonly used power law parametrization, with the particular case of a log-log flat distribution known as “Öpik’s law” [26]. Validity of $f_a \propto a^{-1}$ law up to $a \approx 4600$ AU, which is close to a_{max} of our refined sample of visual binaries, was confirmed by Popova et al [27] and Vereshchagin et al [35] who analyzed the data in the amended 7th Catalog of Spectroscopic Binaries [19] and IDS, respectively. Poveda et al [28], found that Öpik's distribution matches with high degree of confidence binaries with $a \lesssim 3500$ AU (but we note, that selection effects which hamper discovery of the widest systems were not considered, contrary to the abovementioned studies). We also stress, after Heacox [13], that Gaussian distribution of separations encountered in the literature (e. g., Duquennoy & Mayor 1991 [6], Raghavan et al 2010 [30]) is an artefact of data representation. Like Poveda et al [28], we reject Gaussian distribution of stellar separations, since it is hard to envision currently a star-formation process leading to such a distribution.

As for the eccentricity distribution, from physical point of view, one usually prefers in theoretical simulations the “thermal” law $f(e) \sim 2e$ (Ambartsumian 1937) [2], though in observational datasets one finds, e. g., that the eccentricity distribution of wide binaries contains more orbits with $e < 0.2$ and less orbits with $e > 0.8$ (Tokovinin & Kiyaeva 2016 [34]) or a flat distribution in the $e = [0.0 \dots 0.6]$ range and declining one for larger e [30].

Having in mind the difficulties hampering determination of eccentricities from observations and numerous selection effects, we probe three quite different model distributions: “thermal”, flat, and single valued with $e = 0$ for all stars.

The very selection of fundamental parameters for initial distribution is arguable. For instance, primary and secondary masses were considered as fundamental parameters for MS binaries by Malkov [23] and pre-MS binaries by Malkov and Zinnecker [22], while Goodwin [11] has argued that system mass is the more fundamental physical parameter to use. We do not reject possibility to choose and investigate other parameters as fundamental ones in the course of further work.

5 Results and discussion

5.1 Star formation function

Comparison of our simulations with observational data allows us to make the following preliminary conclusions on initial distributions.

Even before application of statistical tests, we should meet a strong and evidently important criterion of validity of the tested combination of initial distributions, namely the agreement between the number of binaries in the simulated datasets and the observed number of visual pairs. This number depends on initial distributions of fundamental variables and changes between 0 and about 15000; an exception is distribution over e which affects the volume of simulated datasets only mildly. Thus, if our accepted normalization [4], along with other used assumptions regarding spatial distribution of visual binaries in solar vicinity is valid, we can exclude certain combinations of initial distributions, based purely on the number of binaries in synthetic dataset. However, our present observational dataset volume (1089 pairs) is limited to binaries having MK spectral classification. Thus, we are careful and do not rely exclusively on this criterion because we allow certain freedom due to simplifications and possible incomplete account of selection effects while constructing the refined observational dataset, as well as to vagueness of theoretical notions on solar vicinity population. This is why we do take into account both number and two sample χ^2 criteria. Nevertheless, one can definitely reject those combinations of initial distributions that lead to the number of binary stars in a synthetic observational dataset significantly less than 1000 (taking present dataset volume $N_{obs} - \sqrt{N_{obs}} \approx 1056$ as lower limit).

Figure 1 represents how the resulting χ^2 statistics are distributed versus number of pairs in the synthetic datasets. The results do not allow us to select “the best” initial distributions over every parameter, but rather to prefer some combinations of initial distributions to others. One may see that no combination leading to acceptable number of pairs in synthetic dataset would give acceptable distribution over angular distances, while magnitude difference and, in some cases, distribution over primary magnitudes, are reproduced better for the same initial conditions. Below there are some figures providing examples of how the same distribution over certain parameter, in different combinations with other initial distributions, leads to better or worse agreement with the observational dataset.

Figure 2 represents an example of how different combinations of initial distributions change resulting synthetic datasets and their agreement with observational one. Four figures demonstrate, in turn, which values of N_{synth} , χ^2 correspond to different initial scenarios (s_0 , s_2 , s_4 , s_5 , see Table 1, Table 2), IMFs (m_0 , m_1), mass ratio initial distribution (q_0 , q_4 , q_5 , applicable solely for the s_2 , s_3 scenarios), and distribution over semi-major axes a_0 , a_1 , a_2 . Scenarios s_0 and s_5 do not involve independent distribution over q ; it is generated as an outcome of the pairing function and IMF, this is why the q -panel contains less dots than the other ones.

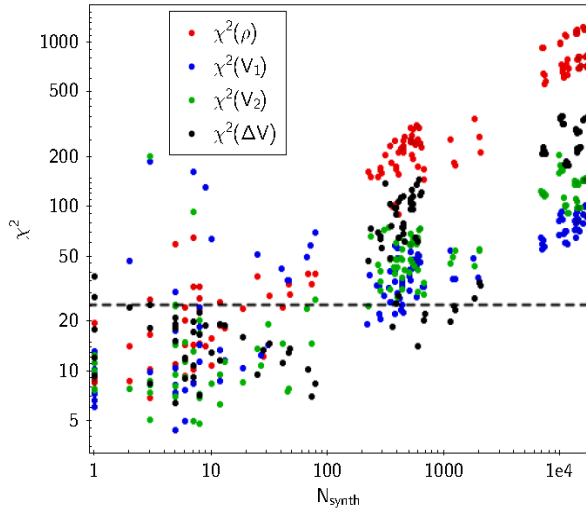


Figure 1 Distribution of resulting χ^2 statistics over number of pairs in the synthetic dataset. Every set of initial distributions of the 144 processed ones results in 4 dots of different colour in this plot. The dashed line marks 5% confidence level of the null-hypothesis (the dots over it correspond to the sets of initial distributions that are rejected at the level of 95%, based on the used observational sample).

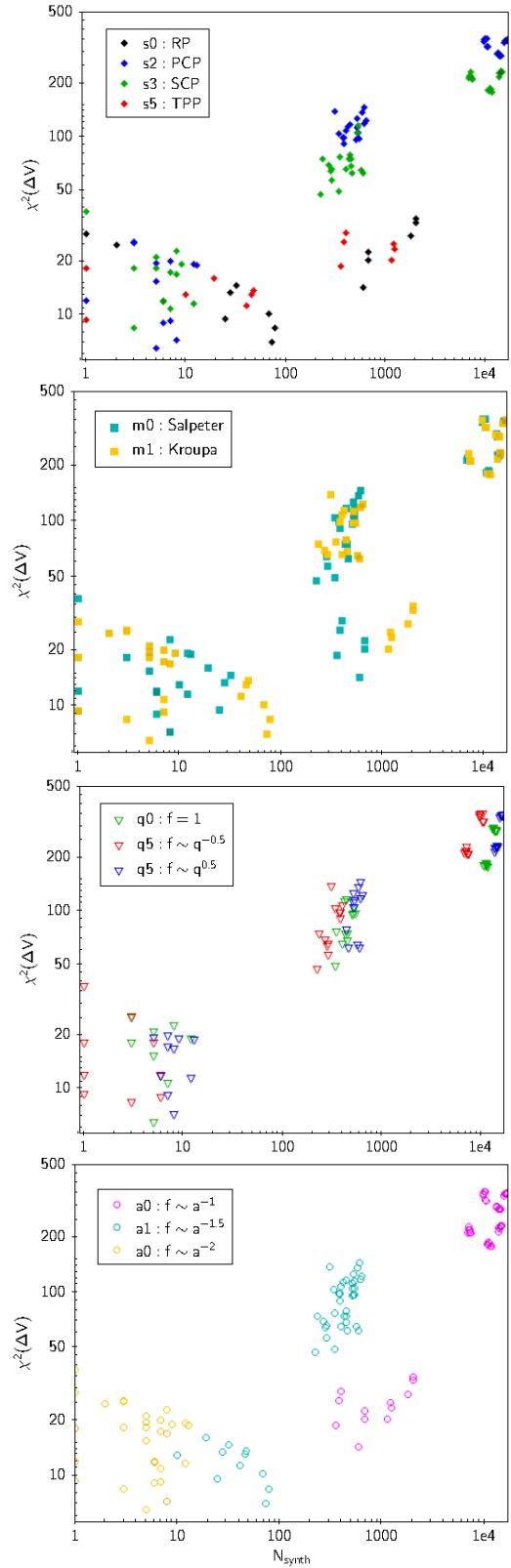


Figure 2 Distribution of resulting χ^2 statistics for magnitude difference ΔV vs. number of pairs in the synthetic datasets, depending on various initial distributions, from top to bottom: pairing scenarios (see Tables 1, Table 2), IMFs, distributions over mass ratio (applicable solely for scenarios s_2 , s_3), and semi-major axes.

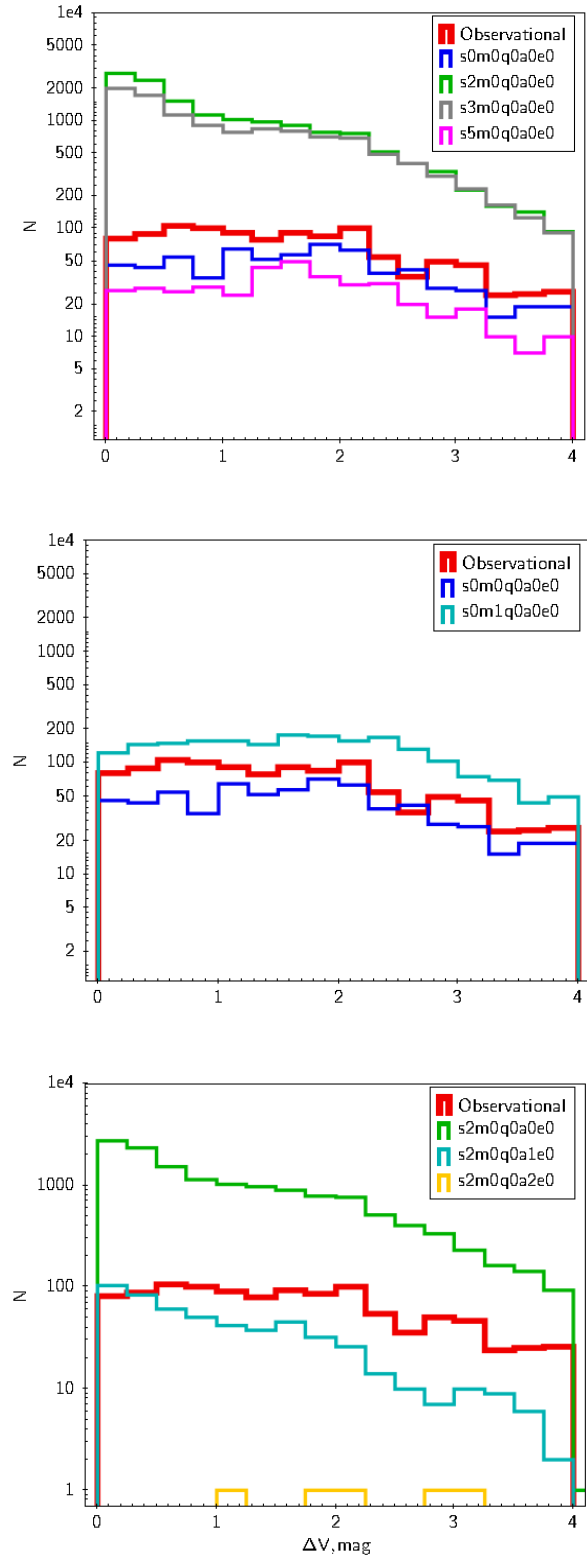


Figure 3 Distributions of resulting synthetic datasets over magnitude difference ΔV for the combinations of initial distributions differing only in (top to bottom) pairing scenarios (see Table 1, Table 2), IMFs, and semi-major axes. The distribution over ΔV for the observational dataset plotted by the bold red line serves as a benchmark.

Figure 3 shows how the distribution over observational parameter magnitude difference changes with the change of one initial distribution (pairing scenarios, IMF, distribution over semi-major axes). The distribution over ΔV for the observational dataset serves as a benchmark.

Based on combination of the two (number and statistical) criteria, we may (very preliminary) state the following.

For the considered observational dataset, RP and TPP pairing scenarios, $s0$ and $s5$ (see Table 1, Table 2), respectively, produce a group of results that seems acceptable in respect of the number of “observed” binaries in the synthetic dataset and, simultaneously, leads to acceptable χ^2 values at least for two observable distributions (V_1 and ΔV).

None of the probed combinations of initial distributions can reproduce observational distribution over angular distance between components adequately (see Figure 1). The cause may lay either with selection effects, that still remain unaccounted for (and then the reconsideration of observational sample is necessary), or in need of other initial distributions.

Kroupa and Salpeter IMF's lead to different number of pairs in the synthetic dataset, however, neither this difference nor χ^2 statistics allows definite choice between them. Kroupa IMF looks slightly more promising, than Salpeter one, however, more accurate conclusion should be postponed, as these two IMF differ actually only in the low-mass region, and the majority of visual binaries in our observational dataset presumably have masses around 1 to $3 M_{\odot}$. The comparison in low-mass region is needed here.

Also, we can not make definite conclusion on the mass ratio q distribution. The q -distributions that we have analyzed in the present study show significant difference only in the low- q region (below $q < 0.5$). In the compilative sample of visual binaries used to construct our benchmark dataset, however, binaries with large magnitude differences (and, thus, low q) are severely underrepresented. This is why we limit refined observational sample so that pairs with low q are excluded. For this reason we can not come to a definite conclusions concerning selection of q -distribution based on this observational sample.

As to the semi-major axes (a) distribution, we have found that power law functions steeper than $a^{-1.5}$ can be excluded from further consideration. Figures 2 and 3 demonstrate that initial distribution $a2$ ($f \sim a^{-2}$, Table 2) leads to inappropriately low volume of synthetic dataset.

It was found also that eccentricity distribution does not influence significantly the resulting distributions.

5.2 Simulation of visibility of binary stars

Depending on the brightness of components and projected separation ρ between them, binary star can be observed as two, one or no source of light, i.e., a part of

binaries can appear as single stars or remain invisible at all. We involve in our simulations the following observational states: “both observed”, “primary only”, “secondary only”, “photometrically unresolved”, and “invisible”. To estimate fraction of simulated pairs, which fall into listed states, we take 0.1 arcsec as a minimum limit for ρ (the limiting value is selected based on analysis of the WDS catalogue), and vary limiting magnitude V_{lim} . We consider a pair to be invisible if its total brightness magnitude exceeds V_{lim} , and to be photometrically unresolved if its ρ does not exceed 0.1 arcsec. We do not pose any restriction to the component magnitude difference. Then, comparing primary and secondary magnitude with V_{lim} , we decide, both or only one component can be observed.

Results of our simulation show that the fraction of photometrically unresolved binaries depends neither on V_{lim} , nor on initial distributions over M , q and e . However, it severely depends on the initial a -distribution: the ratio of unresolved binaries to all visible (as two or one source of light) binary stars equals to about 0.59 ± 0.01 and 0.967 ± 0.003 for $f_a \propto a^{-1}$ and $f_a \propto a^{-1.5}$, respectively.

Fraction of simulated pairs, visible as two sources of lights, hereafter F_{PS} , strongly depends both on a -distribution and V_{lim} . For $f_a \propto a^{-1}$, F_{PS} (depending on q , m and e distributions) varies from 0.01 to 0.19 for $V_{lim} = 16^m$ and from 0.04 to 0.26 for $V_{lim} = 20^m$. F_{PS} values are about ten times lower for $f_a \propto a^{-1.5}$.

Finally, the fraction of simulated stars observed as a single source of light, depends on the F_{PS} as follows: $0.4 - F_{PS}$ for $f_a \propto a^{-1}$ and $0.03 - 0.7 \times F_{PS}$ for $f_a \propto a^{-1.5}$, with no significant dependence on other parameters.

We should note that simple compatibility of synthetic data from initial distributions with observational data is not an ultimate evidence of an adequate modelling, since the observational data are far from being comprehensive. However, the initial distributions we use in our simulations are obtained, checked and used by many other authors, and this suggests that our conclusions are fairly reliable.

6 Future plans

We presented here very preliminary results proving that we need more thorough investigation of the models and comparison with observations to accomplish the task.

To make confident conclusions on BF of binary stars, we need to make comparison of our simulations with other sets of observational data for wide binaries, to cover wider regions of stellar parameters. Also, one of the important further steps is to extend our study to closer binary systems. We will involve basic ideas on evolution of interacting binaries in our simulations, and, consequently, will take into consideration other types of binaries for comparison. It will allow us to make more definite conclusions on BF of more massive stars (as, simulating visual binaries, we deal mostly with

moderate-mass stars), to consider more distant objects, and to involve final stages of stellar evolution into consideration. Having a number of Monte-Carlo simulations representing various observational datasets, we should be able to check if the approximate formula (1) needs reconsideration of remains valid.

Besides the χ^2 two sample test, we plan to consider other statistical methods (e.g., Kolmogorov-Smirnov two sample test) for more reliable interpretation of comparison of our simulation results with observations.

Finally, we aim to consider other parameters as fundamental for initial distributions, e.g., total mass of the binary, angular momentum of a pair, and so on.

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